# **Building Blocks of Population Genetics**Jack Dekkers

## Genotype frequencies, Allele frequencies, Haplotype frequencies, Linkage Disequilibrium, Linkage

<u>Single locus</u> → allele (or gene) frequencies → genotype frequencies

Consider a single locus in a random mating outbred population. The locus has alleles  $A_1$  and  $A_2$  with allele (or gene) frequencies p and q

Under random mating (Hardy Weinberg Equilibrum), the allele received from one parent is *independent* of the allele received from the other parent, resulting in the following relationship between allele and genotype frequencies:

Table 1: Genotype probabilities, single locus two-allele case

. =	Matern		
Paternal allele	$Pr(A_1) = p$	$Pr(A_2) = q$	Marginal prob
$Pr(A_1) = p$	$p^2$	pq	$p^2 + pq = p(p+q) = p$
$Pr(A_2) = q$	pq	$q^2$	$pq + q^2 = q(p+q) = q$
	$p^2 + pq =$	$pq + q^2 =$	
Marginal prob.	p(p+q)=p	q(p+q)=q	

This results in the HWE genotype frequencies:  $p^2$ , 2pq,  $q^2$ 

With <u>multiple loci</u> we also need to consider *haplotypes* and their frequencies, and relationships between *allele*, *haplotype*, and *genotype* frequencies.

**Haplotype** = the combination of alleles at >1 locus that an individual inherited from a parent E.g. an individual with (unordered) genotype  $A_1A_2$  and  $B_1B_2$  at loci A and B, can have the following combinations of haplotype pairs (separated by /):

 $A_1B_1/A_2B_2$   $\rightarrow$  alleles  $A_1$  and  $B_1$  received from one parent and  $A_2$  and  $B_2$  from the other  $A_1B_2/A_1B_2$   $\rightarrow$  alleles  $A_1$  and  $B_2$  received from one parent and  $A_2$  and  $B_1$  from the other

Haplotype frequency = frequency of a given haplotype in a population

With two loci with two alleles, there are 4 possible haplotypes, 16 ordered genotypes (ordered based on haplotypes), and 9 unordered genotypes (see tables 2,3)

Table 2: Haplotype frequencies and genotype frequencies under random mating (HWE)

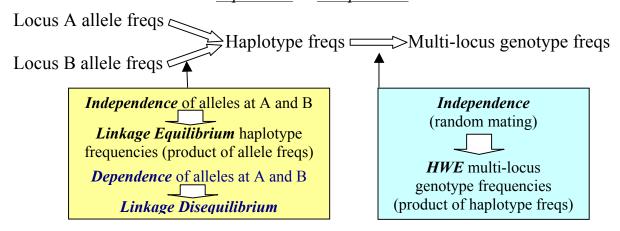
				Mat	ernal	haplotype			
Нар	olotype - freq	$A_1B_1$	r	$A_1B_2$	S	$A_2B_1$	t	$A_2B_2$	u
- 0)	$A_1B_1$	$A_1B_1/A_1B$	$r^2$	$A_1B_1/A_1B_2$	rs	$A_1B_1/A_2B_1$	rt	$A_1B_1/A_2B_2$	ru
Paternal haplotype	$A_1B_2$	$A_1B_2/A_1B$	ı sr	$A_1B_2/A_1B_2$	$s^2$	$A_1B_2/A_2B_1$	st	$A_1B_2/A_2B_2$	su
Pate vaple	$A_2B_1$ t	$A_2B_1/A_1B$	ı tr	$A_2B_1/A_1B_2$	ts	$A_2B_1/A_2B_1$	$t^2$	$A_2B_1/A_2B_2$	tu
,1	$A_2B_2$ $i$	$A_2B_2/A_1B$	<sub>1</sub> ur	$A_2B_2/A_1B_2$	us	$A_2B_2/A_2B_1$	ut	$A_2B_2/A_2B_2$	$u^2$

Table 3: Unordered and ordered genotypes and their frequencies under random mating

Unordered	Frequency	Possible		red genotype		their freque	ncie	es (from Table 1)	)
genotypes	=sum of ordered frequencies	'ordered' ba	ised on	parental origin	ı (pate	rnal haplotype	/mat	ernal haplotyp	e)
$A_1A_1B_1B_1$	$r^2$	$A_{I}B_{I}/A_{I}B_{I}$	$r^2$						
$A_1A_1B_1B_2$	2rs	$A_1B_1/A_1B_2$	rs	$A_1B_2/A_1B_1$	sr				
$A_1A_1B_2B_2$	$s^2$	$A_1B_2/A_1B_2$	$s^2$						
$A_1A_2B_1B_1$	2rt	$A_1B_1/A_2B_1$	rt	$A_2B_1/A_1B_1$	tr				
$A_1A_2B_1B_2$	2ru+2st	$A_1B_1/A_2B_2$	ru	$A_1B_2/A_2B_1$	st	$A_2B_1/A_1B_2$	ts	$A_2B_2/A_1B_1$	ur
$A_1A_2B_2B_2$	2su	$A_1B_2/A_2B_2$	su	$A_2B_2/A_1B_2$	us				
$A_2A_2B_1B_1$	$t^2$	$A_2B_1/A_2B_1$	$t^2$						
$A_2A_2B_1B_2$	2tu	$A_2B_1/A_2B_2$	tu	$A_2B_2/A_2B_1$	ut				
$A_2A_2B_2B_2$	$u^2$	$A_2B_2/A_2B_2$	$u^2$						

The unordered genotype is what is obtained from genotyping, i.e. the genotype at each locus

What is the relationship between haplotype frequencies and the frequencies of alleles that make up each haplotype? This depends on whether the alleles at the two loci are dependent or independent:



#### **Haplotype probabilities** (= frequencies) - two-allele case:

What is the probability of a progeny to receive from a parents: allele  $A_i$  at locus A and allele  $B_i$  at locus B?

## i) if the alleles at the two loci are independent from each other

→ joint probability = product of marginal probabilities

Locus B	Locus A – all	ele frequencies	
allele freq's	$Pr(A_1) = p_A$	$Pr(A_2) = q_A$	Marginal prob
$Pr(B_1) = p_B$	$\Pr(A_I B_I) = p_A p_B$	$\Pr(A_2B_I) = q_{\rm A}p_{\rm B}$	$p_{A}p_{B} + q_{A}p_{B}$ $= p_{B}(p_{A} + q_{A}) = p_{B}$
$Pr(B_2) = q_B$	$\Pr(A_1B_2) = p_{A}q_{B}$	$\Pr(A_2B_2) = q_Aq_B$	$p_{\mathrm{A}}q_{\mathrm{B}}+q_{\mathrm{A}}q_{\mathrm{B}} \ =q_{\mathrm{B}}\left(p_{\mathrm{A}}+q_{\mathrm{A}} ight)=q_{\mathrm{B}}$
Marginal prob	$p_{A}p_{B} + p_{A}q_{B}$ $= p_{A}(p_{B} + q_{B}) = p_{A}$	$q_{A}p_{B} + q_{A}q_{B}$ $= q_{A}(p_{B} + q_{B}) = q_{A}$	

#### ii) What if the alleles at the two loci are NOT independent?

 $\rightarrow$  joint probabilities deviate from product of marginal probabilities (by  $\pm D$ )

	Loca	us A	
Locus B	$Pr(A_1) = p_A$	$Pr(A_2) = q_A$	Marginal prob
$Pr(B_1) = p_B$	$Pr(A_1B_1) = r$ $= p_Ap_B + D$	$Pr(A_2B_1) = t$ $= q_A p_B - D$	$p_{A}p_{B}+D+q_{A}p_{B}-D$ $=p_{B}(p_{A}+q_{A})=p_{B}$
$Pr(B_2) = q_B$	$Pr(A_1B_2) = s$ $= p_Aq_B - D$	$Pr(A_2B_2) = u$ $= q_Aq_B + D$	$p_{A}q_{B} - D + q_{A}q_{B} + D$ $= q_{B} (p_{A} + q_{A}) = q_{B}$
Marginal prob	$p_{A}p_{B} + D + p_{A}q_{B} - D$ = $p_{A}(p_{B} + q_{B}) = p_{A}$	$q_{A}p_{B} - D + q_{A}q_{B} + D$ $= q_{A}(p_{B} + q_{B}) = q_{A}$	

## If alleles are dependent → loci are in Linkage Disequilibrium or in Gametic Phase Disequilibrium

The term 'linkage' in Linkage disequilibrium is actually not quite correct and a bit misleading because disequilibrium can occur between <u>unlinked</u> loci, although it is more likely to be present (and persist) between linked loci (see later). Thus, 'Gametic phase' disequilibrium is a better term; gametic phase refers to the haploid phase of chromosomes and disequilibrium refers to dependence between alleles that make up the haplotypes that are present in the current generation and which originated from the haploid gametes produced by their parents.

$$D = \text{measure of disequilibrium} = D = r - p_A p_B$$

$$Pr(A_1 B_1) - Pr(A_1) Pr(B_1)$$

The value obtained for |D| is the same irrespective of the haplotype used.

$$D = \frac{1}{2} \left[ \Pr\left(\frac{A_1 B_1}{A_2 B_2}\right) - \Pr\left(\frac{A_1 B_2}{A_2 B_1}\right) \right] = ru - st$$
Coupling
heterozygote

Repulsion
heterozygote

#### Other measures of LD:

**D'** = D standardized to make it less dependent on allele frequencies

$$D' = D/D_{max}$$
 where  $D_{max} = \text{Min}(p_A p_B, q_A q_B)$  if  $D < 0$ 

$$D_{max} = \text{Min}(p_A q_B, q_A p_B)$$
 if  $D > 0$ 

 $r^2$  = squared correlation between allele at locus A and allele at locus B- also measures ability  $(R^2)$  to predict allele at locus A from allele at locus B

To derive  $r^2$ : Let X = 1 when allele  $A_1$  present, X = 0 if  $A_2$  present (= Bernoulli var.) Y = 1 when allele  $B_1$  present, Y = 0 if  $B_2$  present (= Bernoulli var.)

$$X = 0$$
 if  $A_2$  present (= Bernoulli var.)  
 $Y = 0$  if  $B_2$  present (= Bernoulli var.)

Then: 
$$cov(X, Y) = E(XY) - E(X) E(Y)$$
  
=  $r - p_A p_B = D$ 

$$\Rightarrow \text{Corr.} = \mathbf{r}_{XY} = \frac{\text{cov}(X, Y)}{\sqrt{\text{var}(X) \text{var}(Y)}} = \frac{D}{\sqrt{(p_A q_A)(p_B q_B)}}$$

	$A_1 X=1$	$A_2 X=0$
$B_1$ $Y=1$	$Pr(A_1B_1) = r$ $XY = 1$	$Pr(A_2B_1) = t$ $XY = 0$
$B_2$ $Y=0$	$Pr(A_1B_2) = s$ $XY = 0$	$Pr(A_2B_2) = u$ $XY = 0$

$$r^2 = r_{XY}^2 = \frac{D^2}{p_A q_A p_B q_B}$$
 (Note: this **r** is different than **r** in the table above)

|D'| and  $r^2$  range between 0 and 1

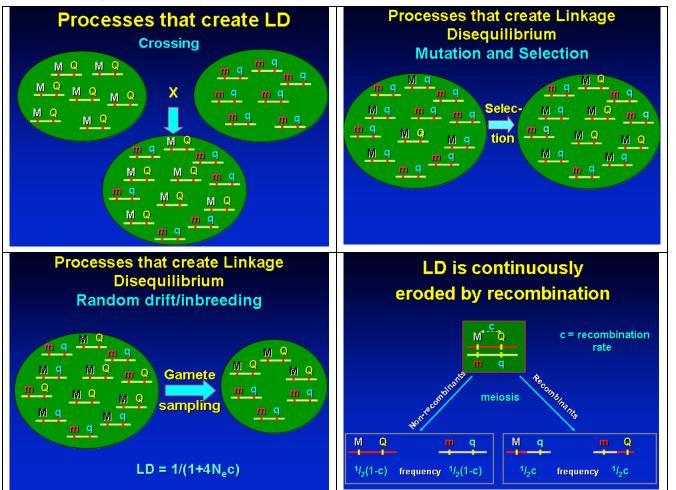
|D'| is strongly inflated if one haplotype has a very low frequency

 $r^2$  is the preferred measure of LD for most uses

## Mechanisms that generate Linkage Disequilibrium (LD)

A variety of mechanisms generate linkage disequilibrium, and several of these can operate simultaneously. They can be separated into:

- 1. **Recurrent factors** operate to create LD each generation
  - a. **Drift** (inbreeding) in small populations by chance or sampling, haplotypes passed on to the next generation are not in LE frequencies
  - b. **Recurrent migration** continuous mixing of populations in which haplotypes occur in different frequencies (e.g.  $Pr(A_1B_1)=1$  for pop. 1 and =0 for pop. 2)
  - c. **Selection** certain haplotypes may be selected upon and increase in frequency
    - selection creates LD between loci that are selected upon (= Bulmer effect)
    - selection with epistasis (certain combinations of alleles are favorable)
       also creates LD between loci involved.
- 2. **Punctual factors** operate only sporadically over time to create LD
  - a. *Mutation* occurs in a specific haplotype, which is then the only haplotype that contains that mutation, resulting it to be in LD with the mutation.
  - b. *One-time admixture/migration/crossing* (e.g. producing  $F_1/F_2$ ) results in mixing populations with different haplotype frequencies
  - c. Population bottleneck / founder effects severe drift from 1-time sampling effects



## **LD** is continuously eroded by recombination: how does *D* change over time?

Let  $r_0$  = frequency of  $A_1B_1$  haplotypes in generation  $0 \rightarrow D_0 = r_0 - p_Ap_B$ 

What is the frequency of  $A_1B_1$  haplotypes in generation 1?

In the following derivation, we will consider parental origin of haplotypes and will let  $_{\bullet}$  indicate 'any' allele, so  $A_1B_1/A_{\bullet}B_{\bullet}$  indicates an individual that received the  $A_1B_1$  from its father and any haplotype  $(A_1B_1 \text{ or } A_1B_2$ or  $A_2B_1$  or  $A_2B_2$ ) from its mother)

There are four ways that parents from generation 0 can generate gametes that carry the  $A_1B_1$ haplotype and that will produce generation 1:

- 1. non-recombinant  $A_1B_1$  haplotype produced by a  $A_1B_1/A_{\bullet}B_{\bullet}$  parent
- 2. non-recombinant  $A_1B_1$  haplotype produced by a  $A_{\bullet}B_{\bullet}/A_1B_1$  parent
- $A_1B_1$  haplotype produced by a  $A_1B_{\bullet}/A_{\bullet}B_1$  parent 3. recombinant
- $A_1B_1$  haplotype produced by a  $A_{\bullet}B_1/A_1B_{\bullet}$  parent 4. recombinant

For case 1, the frequency of  $A_1B_1/A_{\bullet}B_{\bullet}$  parents is  $r_0$ . The frequency of non-recombinant  $A_1B_1$ haplotypes produced by these parents is  $\frac{1}{2}(1-c)$ . Thus, the frequency of  $A_1B_1$  haplotype produced by  $A_1B_1/A_{\bullet}B_{\bullet}$  parents = Prob(1.) =  $\frac{1}{2}(1-c)r_0$ .

Case 2 results in the same frequency: Prob(2.) =  $\frac{1}{2}(1-c)r_0$ .

For case 3, the frequency of  $A_1B_{\bullet}/A_{\bullet}B_1$  parents is  $p_Ap_B$ . The frequency of recombinant  $A_1B_1$ haplotypes produced by these parents is  $\frac{1}{2}$ c, so the overall frequency is  $\frac{1}{2}cp_Ap_B$ .

Case 4.results in the same frequency: Prob(4.) =  $\frac{1}{2}cp_Ap_B$ ..

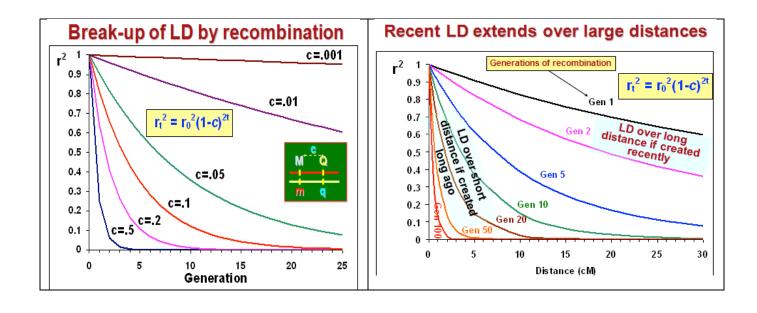
Thus, the overall frequency of  $A_1B_1$  gametes produced by generation 0 is the some of these four mutually exclusive cases:

- $r_1 = r_0(1-c) + p_A p_B c$
- →  $D_1 = r_1 p_A p_B = r_0 (1 c) + p_A p_B c p_A p_B = r_0 (1 c) p_A p_B (1 c) = (r_0 + p_A p_B) (1 c) = D_0 (1 c)$
- →  $D_2 = D_1(1-c) = \{D_0(1-c)\} (1-c) = D_0(1-c)^2$
- →  $D_t = D_0 (1-c)^t$  →  $D_\infty = 0$
- → Erosion of LD by recombination occurs faster when loci are further apart. LD is halved each generation if loci are unlinked ( $c = \frac{1}{2}$ ).

Since  $\mathbf{r}^2 = \frac{D^2}{p_A q_A p_B q_B}$ , LD measured by  $\mathbf{r}^2$  will decline at a rate of  $(1-c)^2$  per

generation:

$$r_t^2 = r_0^2 (1-c)^{2t}$$



## Balance between drift and recombination: in small(er) closed populations

- LD is continuously created by drift (sampling) (small effective population size,  $N_e$ )
- LD is continuously eroded by recombination faster at longer distances

This results in a balance/equilibrium of average LD at a given distance:  $E(r^2_{\infty,c}) = \frac{1}{1 + 4N_e c}$ 

(Sved 1971) Real data  $E(r^2) = 1/(1+4Ne^*c)$ 0.9 commercial chicken breeding line 0.8 Ne = 50 Ne = 100 0.7 Ne = 500 LD (r-squared) Ne = 1000 0.6 Ne = 10000 0.5 ₹ 0.6 0.4 0.3 0.2 0.1 0 0.3 0.4 0.1 0.2 0.5 Genetic distance (Morgans)

Most outbred domesticated plant and animal populations have small(er) (historical) effective population size and drift-recombination balance is expected to be the main contributor to LD → LD expected to be sizeable at short distances, but small at longer distances.

Most human populations have large (historical)  $N_e \rightarrow E(r^2_{\infty,c}) = \frac{1}{1 + 4N_e c}$  is smaller at given distance.

## **Building Blocks of Quantitative Genetics**

#### **Jack Dekkers**

1 locus case:

Genotype  $A_2A_2$   $A_1A_2$   $A_1A_1$ 

Genotypic value  $\mu$ —a  $\mu$   $\mu$ +d  $\mu$ +a  $\mu$  = mid-homozygote value

Single-locus frequencies, genotypic values (deviated from μ), and expectation

T	Frequency,	Genotypic	$Pr(T) \times G_T$
	Pr( <i>T</i> )	value, $G_T$	. ,
$A_1A_1$	$p^2$	а	p²a
$A_1A_2$	2pq	d	2pqd
$A_2A_2$	$q^2$	а	–q²a

**Population mean** =  $\mu + E(G_T) = M = \mu + p^2 a + 2pqd + -q^2 a = \mu + a(p - q) + 2pqd$ 

#### Extension to two loci (without epistasis):

The genotypic value of an individual is the sum of the genotypic values at each locus:

$$G_T = \mu + G_A + G_B$$
  $G_i = \text{genotypic value locus } i$ , as defined for 1-locus case

Now the homozygote "midpoint"  $\mu$  is midway between the best and worst double homozygote (A<sub>1</sub>A<sub>1</sub>B<sub>1</sub>B<sub>1</sub> and A<sub>2</sub>A<sub>2</sub>B<sub>2</sub>B<sub>2</sub>) = (17+3)/2 = 10 in the example below.

Pop. mean = 
$$\mu$$
 + E( $G_T$ ) =  $M$  =  $\mu$  + E( $G_A$  +  $G_B$ ) = E( $G_A$ ) + E( $G_B$ )  
=  $\mu$  + { $a_A(p_A - q_A)$  + 2 $p_Aq_Ad_A$ } + { $a_B(p_B - q_B)$  + 2 $p_Bq_Bd_B$ }

## Spreadsheet Genotypic\_values\_models\_v10.xls

Input parameters:

Homozygo	te midpoir	nt μ =	10
a <sub>A</sub> =	4	a <sub>B</sub> =	3
d <sub>A</sub> =	2	d <sub>B</sub> =	-1
p <sub>A</sub> =	0.6	р <sub>в</sub> =	0.3
q <sub>A</sub> =	0.4	q <sub>B</sub> =	0.7
Linkage D	isequilibri	um D =	0
Recomb	. Rate =	0.2	
Innut mate	iv for opiot	atio offacto	

Input mat	rix for epis	static effects	\$
	$A_1A_1$	$A_1A_2$	$A_2A_2$
B <sub>1</sub> B <sub>1</sub>	0	0	0
$B_1B_1$ $B_1B_2$	0	0	0
$B_2B_2$	0	0	0

Outp						
GE	NOTYPE-BA	ASED M	ODEL FO	OR GENC	TYPIC V	ALUES
	2-locus gen	otypic v	alues	A lo	ocus genot	уре
6	and frequenci	<b>es</b> (random	mating)	$A_1A_1$	$A_1A_2$	$A_2A_2$
		B locus	μ+G <sub>T</sub>	4	2	-4
		genotype	freq	0.36	0.48	0.16
		B <sub>1</sub> B <sub>1</sub>	3	17	15	9
			0.09	0.0324	0.0432	0.0144
		B <sub>1</sub> B <sub>2</sub>	-1	13	11	5
Pop	ulation mean		0.42	0.1512	0.2016	0.0672
	M = 10.14	$B_2B_2$	-3	11	9	3
ne	ew <sup>µ</sup> = 10		0.49	0.1764	0.2352	0.0784
	$\mu_{A} = 8.38$	Average at	A locus	12.38	10.38	4.38
	$\mu_{\rm B}$ = 11.76	Average at	B locus	14.76	10.76	8.76
Re-d	calculated 1-locus	additive, do	minance	+a	d	-a
	and genoty	pic values	$G_A =$	4	2	-4
	with epistas	sis	$G_B =$	3	-1	-3

Extension to many loci:  $G_T = \sum G_i$  summation is over all loci

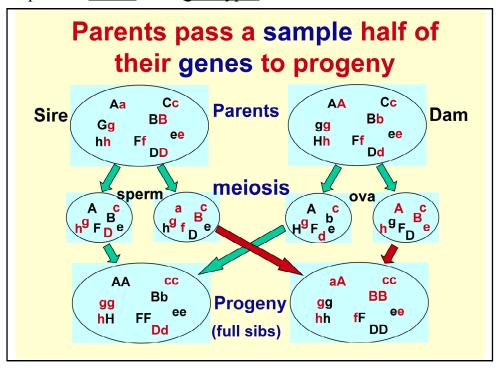
Homozygote "midpoint"  $\mu$  is average of the best and the worst multi-homozygote Population mean =  $\mu$  + E( $G_T$ ) = M =  $\mu$  + E( $\Sigma G_i$ ) =  $\mu$  +  $\Sigma$ E( $G_i$ ) =  $\mu$  +  $\Sigma$ { $a_i(p_i - q_i)$  +  $2p_iq_id_i$ } =  $\Sigma a_i(p_i - q_i)$  +  $2\Sigma p_iq_id_i$ 

## Allele-based models for additive effects

In practice, we are interested in selecting the 'best' individuals to be used as parents to breed the next generation; we want to select individuals whose <u>progeny</u> have the highest expected phenotype, i.e. whose progeny have the highest expected genotypic value.

To identify these individuals we need to know how the genotypic value of progeny relates to the genotypic value of their parents.

Models described in terms of **a** and **d** are for genotypic values for whole genotypes. But individuals pass on alleles NOT genotypes



The breeding value of an individual is defined to quantify an individual's value as a parent. It is related to the <u>expected</u> genotypic or phenotypic values of that individual's progeny.

An individual's **breeding value** =  $2 \times \text{expected deviation of the mean phenotype}$  of an individual's progeny from population mean (M) when mated at random to other individuals from the population.

$$A_i = 2 E(P_{\text{progeny}} - M)$$

In general: an individual's breeding value is the sum of the average effects of the alleles that the individual carries:  $A_{ij} = \alpha_i + \alpha_j$ 

Average effect  $\alpha_i$  = Average deviation from the population mean of individuals who received allele i (i.e.  $A_1$  or  $A_2$ ) from one parent and the other allele at random (i.e.  $A_1$  with freq. p and  $A_2$  with freq. q)

	"Other Allele	e" (=random)	Mean $G_T$	Mean of $G_T$
	$A_1$	$A_2$	of resulting	deviated from
"Allele i"	$Pr(A_1) = p$	$Pr(A_2) = q$	individuals	Population mean
				$\alpha_1 = pa + qd - M$
$A_1$	$G_{A1A1} = a$	$G_{A1A2} = d$	pa + qd	$=q[a+(q-p)d] = q\alpha$
				$\alpha_2 = pd - qa - M$
$A_2$	$G_{A2A1} = d$	$G_{A2A2} = -a$	pd – qa	$=-p[a+(q-p)d]=-p\alpha$

Another important concept/quantity is the average allele substitution effect ( $\alpha$ ):

Average allele substitution effect = 
$$\alpha = \alpha_1 - \alpha_2 = a + (q - p)d$$

= average effect on the genotypic value of substituting a random  $A_2$  allele for an  $A_1$  allele **One locus multiple alleles**: an individual's breeding value is the sum of the average effects of the alleles that the individual carries:  $A_{ij} = \alpha_i + \alpha_j$ 

For two loci: individual with alleles i and j at locus A and alleles k and l at locus B:  $A_{ijkl} = \alpha_{Ai} + \alpha_{Aj} + \alpha_{Bk} + \alpha_{Bl} \quad \text{with each } \alpha_{ni} \text{ derived as above for 1-locus case (no epistasis)}$ 

**Many loci**: 
$$A_{ijkl} = \sum_{locus \ l=1}^{n} \sum_{alle \ le \ i=1}^{2} \alpha_{li}$$

sum average effects over all *n* loci and the individual's two alleles at each locus

## Spreadsheet 'Genotypic\_value\_models.v10.xls'

	Average all	ele effects			Substitution effect	]    ,	Single locus Genotypic and	Breeding V
ocus A	A1 =	1.44	A2 =	-2.16	A = 3.6	6 Т	deviated from population	on mean (M)
ocus B	<sub>B1</sub> =	1.82	B2 =	-0.78	<sub>B</sub> = 2.6			
All values a	are now devi	ated from t	he populati	on mean, N	Л.	4 -		
		ΑI	ocus genot	уре		e l		
		$A_1A_1$	$A_1A_2$	$A_2A_2$	1	value 2 -		
	( +G <sub>T</sub> )-M	2.24	0.24	-5.76		6 -		
	Additive	2.88	-0.72	-4.32		lie 0		
B locus	Dom.dev.	-0.64	0.96	-1.44		reeding		
Genotype	Epistasis					1 0 1		
B <sub>1</sub> B <sub>1</sub>	4.62	6.86	4.86	-1.14		ຈ -2 -		
	3.64	6.52	2.92	-0.68		Υpi		
	0.98	0.34	1.94	-0.46		[ 월 -4 -		
	_	0.00	0.00	0.00		Genotypi	O	
$B_1B_2$	0.62	2.86	0.86	-5.14		-6 -		
	1.04	3.92	0.32	-3.28				
	-0.42	-1.06	0.54	-1.86		_8 _		
	_	0.00	0.00	0.00		-6 -	A la avec C	Diame
$B_2B_2$	-1.38	0.86	-1.14	-7.14			71100000	B locus
	-1.56	1.32	-2.28	-5.88			—O— A locus Br.val. —	O− B locus I
	0.18	-0.46	1.14	-1.26				•
		0.00	0.00	0.00				

#### Alternate derivation of allele substitution effect based on

## Linear regression on number of '1' alleles

Allele substitution effects can also be derived by analyzing phenotype (or the genotypic value) by a linear regression model on the number of '1' alleles that an individual carries (as we have done is some of the homeworks):

Linear regression of Y on X:  $Y = M + \hat{b}_{YX}(X - \overline{X}) + e$  $\hat{Y} = M + \hat{b}_{YY}(X - \overline{X})$ 

In this case, X = # 1 alleles:  $X_T = \{2, 1, 0\}$  for  $T = \{A_1A_1, A_1A_2, A_2A_2\}$ 

Genotype		# A <sub>1</sub>	Frequency		
T	$G_T$	$X_T$	f	$f^*X_T^2$	$F^*G_T^*X_T$
$A_1A_1$	а	2	$p^2$	4p <sup>2</sup>	2p²a
$A_1A_2$	d	1	2pq	2pq	2pqd
$A_2A_2$	-a	0	$q^2$	0	0
SUM				4p <sup>2</sup> +2pq	2p <sup>2</sup> a+2pqd

$$\hat{b} = \frac{\text{cov}(G_T, X_T)}{\text{var}(X_T)} = \frac{E(G_T X_T) - E(G_T)E(X_T)}{E(X_T^2) - E(X_T)^2} = \frac{\left[2p^2a + 2pqd\right] - \left[a(p-q) + 2pqd\right]\left[2p^2 + 2pq\right]}{\left[4p^2 + 2pq\right] - \left[2p^2 + 2pq\right]^2}$$

Interpretation of Eregression coefficient  $\hat{b}$ :

- When  $X_T$  increases by 1,  $\hat{Y}$  increases by  $a + (q p)d = \alpha$  = allele substitution effect
- When  $X_T$  increases by 1, an allele substitution has occurred  $\Rightarrow \hat{b} = \alpha$

## Allele-based models for dominance and epistatic effects

When d is not 0, breeding values will not explain everything about the genotypic value:

Single locus example: p=0.6; a=+4; d=1; M=0  $\Rightarrow \alpha = a + (q - p)d = +3.8$  $\Rightarrow \alpha_1 = q\alpha = 0.4*3.8 = +1.52$ 

 $\rightarrow \alpha_2 = -p\alpha = -0.6*3.8 = -2.28$ 

Genotype	Fre-	Genotypic value (G)	Breeding value	Dominance deviation
T	quency	deviated from M	A	$\delta = G - A$
$A_1A_1$	0.36	+2.72	$2\alpha_1 = +3.04$	-0.32
$A_1A_2$	0.48	-0.28	$\alpha_1 + \alpha_2 = -0.76$	+0.48
$A_2A_2$	0.16	-5.28	$2\alpha_2 = -4.56$	-0.72

These differences between the single-locus genotypic and breeding values are called 'dominance deviations'.

Based on this, the genotypic value at a single locus of an individual that has alleles i and j at that locus (deviated from the population mean, M) can be written as:

$$G_{ij} = \alpha_i + \alpha_j + \delta_{ij} = A_{ij} + \delta_{ij}$$

 $\alpha_i$  = average effect of allele i  $\alpha_j$  = average effect of allele j  $\delta_{ij}$  = dominance deviation effect of the interaction of alleles i and j

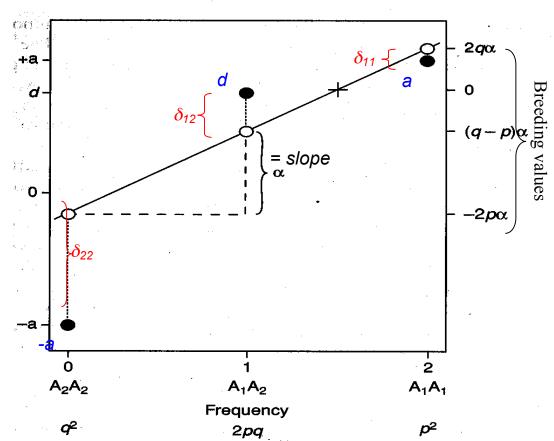
Dominance devations can also be derived as a function of allele frequencies and d

	$G_T$	Average a	Dom.dev.	
Genotype <i>T</i>	(deviated from μ)	$oldsymbol{lpha}_i$	$lpha_{\!\scriptscriptstyle j}$	$\delta_{ij}$
$A_1A_1$	а	$\alpha_1 = q\alpha$	$\alpha_1 = q\alpha$	$-2q^2d$
$A_1A_2$	d	$\alpha_1 = q\alpha$	$\alpha_2$ =- $\rho\alpha$	2pqd
$A_2A_2$	-a	$\alpha_2$ =- $p\alpha$	$\alpha_2$ =- $\rho\alpha$	$-2p^2d$

Graphical representation of average effects and dominance deviations

Dominance deviations are the residuals from the regression of genotypic values on the number of A<sub>1</sub> alleles.

The regression line represents the breeding values



#### Extension to two loci (no epistasis)

and

With two loci, the genotypic value is the sum of the individual's genotypic value at each locus:  $G_T = G_A + G_B$   $G_i = \text{genotypic value locus } i$ , as defined for 1-locus case

And the genotypic value at each locus can be partitioned into additive and dominance effects:

$$G_{Aij} = \alpha_{Ai} + \alpha_{Aj} + \delta_{Aij}$$
 $G_{Bij} = \alpha_{Bi} + \alpha_{Bj} + \delta_{Bij}$ 

Thus the overall genotypic value can be written as the sum of average allele effects and dominance deviations as:

$$G_{T} = \alpha_{Ai} + \alpha_{Aj} + \frac{\delta_{Aij}}{\delta_{Aij}} + \alpha_{Bi} + \alpha_{Bj} + \frac{\delta_{Bij}}{\delta_{Bij}}$$

$$G_{T} = \alpha_{Ai} + \alpha_{Aj} + \alpha_{Bi} + \alpha_{Bj} + \frac{\delta_{Aij}}{\delta_{Bij}} + \frac{\delta_{Bij}}{\delta_{Bij}}$$

The sum of average allele effects define the **breeding value**:  $A_T = \alpha_{Ai} + \alpha_{Aj} + \alpha_{Bi} + \alpha_{Bi}$ 

The sum of dominance deviations define the **dominance effect**:  $D_T = \delta_{Aij} + \delta_{Bij}$ 

Thus the genotypic value can be written as the breeding value and its dominance effect:

$$G_T = A_T + D_T$$

Homozyg	ote midpoir	nt =	10				Αle	ocus genoty	ре
a <sub>A</sub> =	4	a <sub>B</sub> =	3				$A_1A_1$	$A_1A_2$	$A_2A_2$
d <sub>A</sub> =	2	d <sub>B</sub> =				$(+G_T)-M$	2.24	0.24	-5.76
	0.6		0.3			Additive	2.88	-0.72	-4.32
					B locus	Dom.dev.	-0.64	0.96	-1.44
<b>4</b> <sub>A</sub> -	0.4	q <sub>B</sub> =			Genotype	Epistasis			
Linkage I	Disequilibri	um D =	0		B <sub>1</sub> B <sub>1</sub>	4.62	6.86	4.86	-1.14
	Recomb	. Rate =	0.2			3.64	6.52	2.92	-0.68
Input mat	rix for epist	atic effects	•			0.98	0.34	1.94	-0.46
	$A_1A_1$	$A_1A_2$	$A_2A_2$			_	0.00	0.00	0.00
B <sub>1</sub> B <sub>1</sub>	0	0	0		$B_1B_2$	0.62	2.86	0.86	-5.14
$B_1B_2$	0	0	0			1.04	3.92	0.32	-3.28
$B_2B_2$		0	0			-0.42	-1.06	0.54	-1.86
D <sub>2</sub> D <sub>2</sub>	U	U	U			_	0.00	0.00	0.00
_	BASED MO			VALUES	$B_2B_2$	-1.38	0.86	-1.14	-7.14
Locus A	verage allele effe		-2.16	ostitution effect  A = 3.6		-1.56	1.32	-2.28	-5.88
Locus B	<sub>A1</sub> = 1.44		-0.78	<sub>A</sub> = 3.6 <sub>B</sub> = 2.6		0.18	-0.46	1.14	-1.26
All values are	now deviated from					_	0.00	0.00	0.00

Extension to multiple loci (no epistasis):

$$G_T = A_T + D_T$$
  $A_T = \sum_{locus \ l=1}^n \sum_{all \ el \ e \ i=1}^2 \alpha_{li}$   $D_T = \sum_{locus \ l=1}^n \delta_{lij}$ 

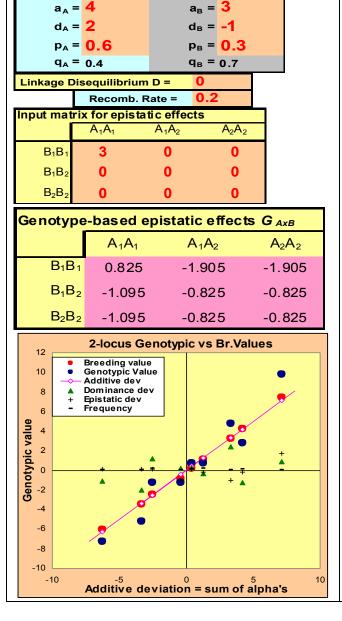
#### **Epistatic Deviations**

Homozygote midpoint

When epistatic effects are present, the genotypic value of an individual can not be written as a simple sum of the genotypic value at each locus but an effect of the interaction between loci needs to be added: For two loci:  $G_T = G_A + G_{B} + G_{A\times B}$ 

Similarly, the genotypic value of an individual can also not be written as the sum of a breeding value and a dominance deviation but an epistatic deviation effect  $(I_T)$  needs to be added:  $G_T = A_T + D_T + I_T$ 

Epistatic deviation effects for each individual (or multi-locus genotype) can be calculated by subtraction, after the additive and dominance deviation effects have been computed as described before:  $I_T = G_T - A_T - D_T$ 



ALLELE-B	SASED MO	DEL FOR G	ENOTYP	IC VALUES
Ave	rage allele effec	ets		Substitution effect
Locus A	<sub>A1</sub> = 1.5048	<sub>A2</sub> = _	2.2572	<sub>A</sub> = 3.762
Locus B	<sub>B1</sub> = 2.0468	<sub>B2</sub> = -	0.8772	<sub>B</sub> = 2.924
		ΑI	ocus ge	notype
		$A_1A_1$	$A_1A_2$	$A_2A_2$
	( +G <sub>T</sub> )-M	2.41	0.14	-5.86
	Additive	3.01	-0.7	-4.51
B locus	Dom.dev.	-0.60	0.90	-1.34
Genotype	Epistasis			
B <sub>1</sub> B <sub>1</sub>	5.60	9.76	4.76	-1.24
	4.09	7.10	3.34	-0.42
	1.51	0.91	2.40	0.17
	_	1.75	-0.98	-0.98
$B_1B_2$	0.52	2.76	0.76	-5.24
	1.17	4.18	0.42	-3.34
	-0.65	-1.24	0.25	-1.99
		-0.17	0.10	0.10
$B_2B_2$	-1.48	0.76	-1.24	4 -7.24
	-1.75	1.26	-2.5	-6.27
	0.28	-0.32	1.17	-1.07
	_	-0.17	0.10	0.10

## **Genetic Variance Components**

## **Single-locus model**

Genotype 
$$A_2A_2$$
  $A_1A_2$   $A_1A_2$  Genotypic value  $-a$   $0$   $d$   $a$ 

T	Fre-	Genotypic	Pr( <i>T</i> )	$(G_T)^2$	Pr( <i>T</i> )	Breeding	Domi-
	quency	value	×		×	value	nance dev.
	Pr( <i>T</i> )	$G_{T}$	$G_{T}$		$(G_T)^2$	$A_T = \alpha_i + \alpha_j$	$D_{\mathcal{T}} = \delta_{ij}$
$A_1A_1$	$p^2$	а	p <sup>2</sup> a	a <sup>2</sup>	$p^2a^2$	$2q\alpha$	$-2q^2d$
$A_1A_2$	2pq	d	2pqd	$d^2$	2pqd <sup>2</sup>	$(q-p)\alpha$	2pqd
$A_2A_2$	$q^2$	–a	$-q^2a$	a <sup>2</sup>	$q^2a^2$	$-2p\alpha$	$-2p^2d$

Genetic model for genotypic values:  $G_T = A_T + D_T$  = Breeding value + Dominance dev.

Variance of genotypic values in a population = (Total) Genetic variance =  $V_G$ 

$$V_G = \text{var}(G_T) = p^2 a^2 + 2pqd^2 + q^2 a^2 - E(G_T)^2 = 2pq[a + (q - p)d]^2 + (2pqd)^2$$
  
Using  $\alpha = a + (q - p)d = \text{allele substitution effect:}$   $V_G = 2pq \alpha^2 + (2pqd)^2$ 

Additive genetic variance = variance of breeding values in a population =  $V_A$ 

$$V_A = \text{var}(A_T) = p^2 (2q\alpha)^2 + 2pq[2(q-p)\alpha]^2 + q^2 (-2p\alpha)^2 - 0^2 = 2pq\alpha^2$$
 (Note that  $E(A_T) = 0$ )

**Dominance variance = variance of Dominance deviations** in a population =  $V_D$  Using the table on p1, the variance of dominance deviations in the population is :

$$V_D = \text{var}(D_T) = \text{var}(\delta_{ii}) = p^2(-2q^2d)^2 + 2pq(2pqd)^2 + q^2(-2p^2d)^2 - 0^2 = (2pqd)^2$$
 (E(D<sub>T</sub>)=0)

Genotypic variance = 
$$V_G = 2pq\alpha^2 + (2pqd)^2 = V_A + V_D$$
  
= Additive Variance + Dominance Variance

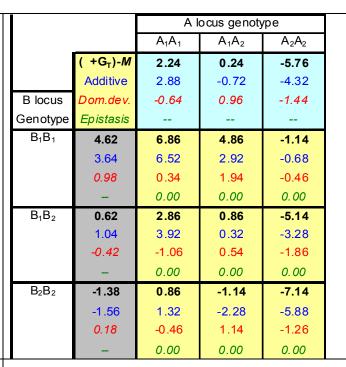
Note:  $COV(A_T,D_T) = 0$ ; i.e. breeding values and dominance deviations are independent

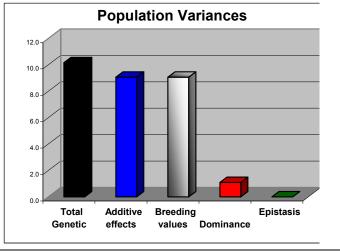
#### **Extension to two loci** – first without epistasis:

Homozyg	ote midpoir	nt =	10			
a <sub>A</sub> =	<b>4</b>	a <sub>B</sub> =	3			
d <sub>A</sub> =	2	d <sub>B</sub> =	-1			
p <sub>A</sub> =	0.6	р <sub>в</sub> =	0.3			
q <sub>A</sub> =	0.4	q <sub>в</sub> =	0.7			
Linkage	Disequilibriu	um D =	0	Ī		
	Recomb	. Rate =	0.2			
Input mat	rix for epist					
	$A_1A_1$	$A_1A_2$	$A_2A_2$			
B <sub>1</sub> B <sub>1</sub>	0	0	0			
B <sub>1</sub> B <sub>2</sub>	0	0	0			
B <sub>2</sub> B <sub>2</sub>	0	0	0			
ALLELE	-BASED N	<b>IODEL</b> F	OR GENO	- TYF	PIC VAL	UES
	Average allele e					ion effect
Locus A	<sub>A1</sub> = 1.44	1	<sub>A2</sub> = -2.16		Α -	= 3.6
Locus B	<sub>B1</sub> = 1.82		<sub>B2</sub> = -0.78			= 2.6
All values a	re now deviated	d from the po	pulation mea	n, M		
Population	n variances	A locus	R locus	Do	nulation	Percent

Population variances	A locus	B locus	Population	Percent
Total Genetic	7.142	3.016	10.158	100.0%
Additive effects	6.221	2.839	9.060	89.2%
Breeding values	6.221	2.839	9.06	
Dominance	0.922	0.176	1.098	10.8%
Epistasis			0.000	0.0%

<sup>&</sup>quot;Additive effects" refer to breeding values computed as the sum of average allele effects. "Breeding values" are computed based on the expected progeny means





Extended to >2 loci, this gives:

$$V_G = \sum V_{Gi} = \sum \{2p_iq_i\alpha_i^2 + (2p_iq_id_i)^2\} = \sum V_{Ai} + \sum V_{Di} = V_A + V_D$$

with: 
$$V_A = \sum V_{Ai} = \sum 2p_iq_i\alpha_i^2$$
 and  $V_D = \sum V_{Di} = \sum (2p_iq_id_i)^2$ 

→ the genetic, additive, and dominance variances for a quantitative trait are the simple sum of the genetic, additive, and dominance variances at each locus that affect the trait.

#### With Epistatic effects = Interactions between the effects that loci have on phenotype

 $G_T = G_A + G_B + G_{AxB}$ Two locus example:

Genotype-based model

$$G_T = \alpha_{Ai} + \alpha_{Aj} + \alpha_{Bi} + \alpha_{Bj} + \delta_{Aij} + \delta_{Bij} + I_{AB}$$
 Allele-based model  
=  $A + D + I$ 

## Epistatic variance = variance of epistatic deviations in a population = $V_I$ = var( $I_{AB}$ )

→ Complete partitioning of genetic variance:  $V_G = V_A + V_D + V_I$  Note: all cov's = 0

$$V_G = V_A + V_D + V_I$$
 Note: all cov's =

 $\rightarrow$  Epistatic variance can be obtained by difference:  $V_I = V_G - V_A - V_D$  see spreadsheet for ex.

Input mat	Input matrix for epistatic effects						
	$A_1A_1$	$A_1A_2$	$A_2A_2$				
B <sub>1</sub> B <sub>1</sub>	3	0	0				
B <sub>1</sub> B <sub>2</sub>	0	0	0				
$B_2B_2$	0	0	0				

Population variance	A locus	B locus	Population	Percent
Total Genetic	7.595	4.009	11.774	100.0%
Additive effects	6.793	3.591	10.384	88.2%
Breeding values	6.793	3.591	10.41357	
Dominance	0.801	0.418	1.220	10.4%
Epistasis			0.170	1.4%

				Αb	ocus	genoty	ype	
			A <sub>1</sub> ,	Α1	Α	1A2		$A_2A_2$
	(μ+G <sub>T</sub> ).	м	2.4	11	0	.14		-5.86
	Additiv	/e	3.0	01	-0	1.75		-4.51
B locus	Dom.de	eν.	-0.0	50 T	0	.90		-1.34
Genotype	Epistas	sis		-				
B <sub>1</sub> B <sub>1</sub>	5.60		9.7	76	4	.76		-1.24
	4.09		7.1	10	3	.34		-0.42
	1.51		0.9	91	2	.40		0.17
			1.7	75	-0	.98		-0.98
B <sub>1</sub> B <sub>2</sub>	0.52		2.7	76	0	.76		-5.24
	1.17		4.1	18	0	.42		-3.34
	-0.65		-1.3	24	0	.25		-1.99
			-0.1	-0.17		.10		0.10
B <sub>2</sub> B <sub>2</sub>	-1.48		0.7	.76		-1.24		-7.24
	-1.75		1.2	26	-2	.51		-6.27
	0.28		-0.3	32	1	.17		-1.07
			-0.1	17	0	.10		0.10
Population	means	Α	locus	B lo	cus	Populat	ion	Populatio
Genotypic v	alues		0.00	0.0	00	0.00		<mark>means ar</mark>
Additive			0.00	0.0	00	0.00	1	always
Dominance	deviations	(	0.00	0.0	00	0.00		equal to
Fuintania da						0.00		

Population Variances
12.0
Total Additive Breeding Epistasis Genetic effects values Dominance

Covariances	Addit.	Dom.dev	Covariances
Dom.dev	0.00		are always
Epist.dev	0.00	0.00	<mark>zero</mark>

Additive, dominance, and epistatic effects are independent (no covariances)

In a typical population, most genetic variance is additive — see also Hill et al. PLOS Genetics (2008)

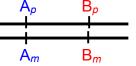
## Impact of Linkage disequilibrium on genetic variances

Consider 2 linked loci, A and B:

Paternal gamete *p* 

Each individual has a paternal and a maternal gamete

Maternal gamete *m* 



The breeding value is the sum of the average effects of the paternal and maternal alleles:

$$A = \alpha_{Ap} + \alpha_{Bp} + \alpha_{Am} + \alpha_{Bm}$$

If the loci are in LD  $\rightarrow$  allele states (0/1) at two loci on the same gamete are not independent

 $\rightarrow$  they have a non-zero covariance ( $r^2$  = squared correl. of allele states (0/1) is > 0)

Then, the variance caused by the additive effect of the paternal (or maternal) gamete is

$$\operatorname{var}(\alpha_{\mathsf{A}p} + \alpha_{\mathsf{B}p}) = \operatorname{var}(\alpha_{\mathsf{A}p}) + \operatorname{var}(\alpha_{\mathsf{B}p}) + 2\operatorname{cov}(\alpha_{\mathsf{A}p}, \alpha_{\mathsf{B}p})$$

$$= \frac{1}{2}V_{\mathsf{A}_{\mathsf{A}}} + \frac{1}{2}V_{\mathsf{A}_{\mathsf{B}}} + 2\mathsf{D}_{\mathsf{A}\mathsf{B}}\alpha_{\mathsf{A}p}\alpha_{\mathsf{B}p}$$

$$\operatorname{where } \mathsf{D}_{\mathsf{A}\mathsf{B}} = \operatorname{LD} \mathsf{A}, \mathsf{B}$$

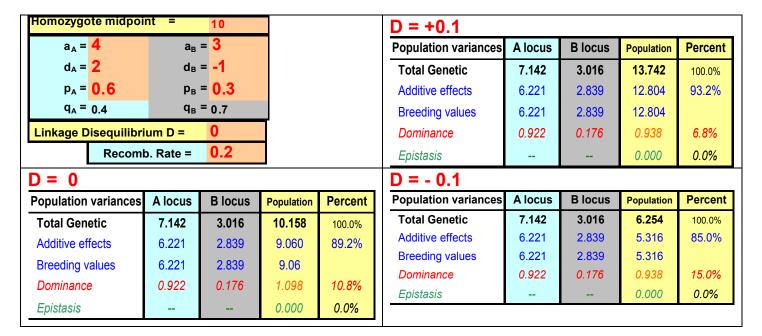
→ Additive genetic variance =  $V_A = V_{A_A} + V_{A_B} + 4D_{AB}\alpha_{Ap}\alpha_{Bp}$ 

see spreadsheet for example

→ Dominance genetic variance =

$$V_D = \text{var}(\delta_{Apm} + \delta_{Bpm}) = \text{var}(\delta_{Apm}) + \text{var}(\delta_{Bpm}) + 2\text{cov}(\delta_{Apm}, \delta_{Bpm})$$
$$= V_{D_A} + V_{D_B} + 8D_{AB}^2 d_A d_B$$

(D<sup>2</sup> because dominance is based on combinations of paternal and maternal alleles)



Note that individual locus variances are not affected by LD but across locus variances are (because of non-zero covariances).

Whether  $V_A$  increases or decreases depends on whether the favorable alleles are in repulsion or coupling phase

Whether  $V_D$  increases or decreases depends on whether d has the same sign for both loci.

## Phenotypic and Environmental Effects and Variances

#### **Models for Environmental Effects**

Phenotype for a quantitative trait is determined by genetic and environmental factors:

$$P = \mu + G + E$$

u includes the mean and systematic (environmental) effects

G = genotypic value

E = Random environmental effects

#### Partitioning of phenotypic variance

Phenotypic variance = var. of phenotypes in a pop. <u>after removal/adjustment for syst.</u> effects

$$= V_P = \text{var}(P-\mu) = \text{var}(G+E) = \text{var}(G) + \text{var}(E) + 2\text{cov}(G,E)$$

If genotypes are distributed at random relative to random environmental effects  $\rightarrow \text{cov}(G,E)=0$ 

$$\rightarrow$$
  $V_P = V_G + V_E = V_A + V_D + V_I + V_E$ 

## Relative importance of the genetic component

Genetic variance as a fraction of the phenotypic variance:

Broad sense heritability  $H^2 = \frac{V_G}{V_P}$  = proportion of phen. var. in a pop. that is genetic

Narrow sense heritability  $h^2 = \frac{V_A}{V_P}$  = proportion of phen. var. that is <u>additive</u> genetic

#### **MODELS FOR TRAITS WITH REPEATED MEASURES**

**General** environmental effects = Eg =Effects that are common to each measurement

**Special** environmental effects = Es = Effects that are specific to a given measurement

$$P_{ij} = G_i + Eg_i + Es_{ij}$$
  $P_{ij} = j^{th}$  measurement of phenotype on  $i^{th}$  individual  $G_i$  and  $Eg_i$  are common to all measurements on individual  $i$   $Es_{ij} =$  special envir. effect for  $j^{th}$  measurement on  $i^{th}$  individual

This also allows random environmental variance to be separated into variances due to General versus Special environmental effects:

$$V_P = V_G + V_{Eg} + V_{Es}$$
 Note:  $cov(Eg,Es)=0$ 

**Repeatability** = r = correlation between repeated measures on the same individual (Assume that  $V_{Es}$  and therefore  $V_P$  is the same for each measurement)

$$r = \text{corr}(P_{ij}, P_{ik}) = \frac{\text{cov}(P_{ij}, P_{ik})}{\sqrt{\text{var}(P_{ij})\text{var}(P_{ik})}} = \frac{\text{cov}(P_{ij}, P_{ik})}{\sqrt{V_{P}V_{P}}} = \frac{\text{cov}(P_{ij}, P_{ik})}{V_{P}}$$

$$cov(P_{ij}, P_{ik}) = cov(G_i + Eg_i + Es_{ij}, G_i + Eg_i + Es_{ik})$$

$$= cov(G_i, G_i) + cov(Eg_i, Eg_i) + Cov(Es_{ij}, Es_{ik})$$

$$= V_G + V_{Eg} + 0$$
 special env. effects are independent

→ repeatability =  $r = (V_G + V_{Eg}) / V_P$  = prop. of  $V_P$  that is due to effects that are consistent across measurements (G + Eg)

$$1 - r = V_{Es} / V_P$$
 = prop. of  $V_P$  that is due to effects that differ between measurements (Es)

#### **CORRELATED TRAITS**

Phenotypic correlation = correlation between phenotypes on traits 1 and 2 on same individual = caused by genetics and environment

 $P_1 = A_1 + D_1 + I_1 + E_1$ 

 $P_2 = A_2 + D_2 + I_2 + E_2$ 

#### Phenotype trait 1

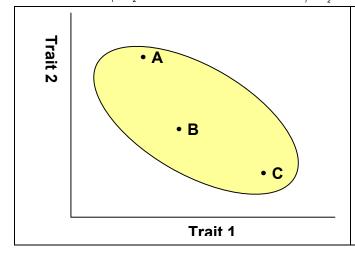
(Additive) genetic correlation =  $\mathbf{r_A}$  Environmental correlation =  $\mathbf{r_E}$ 

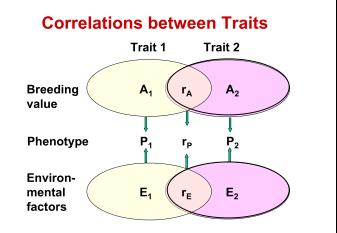
#### Phenotype trait 2

$$r_P = \frac{Cov(P_1, P_2)}{\sigma_{P_1} \sigma_{P_2}}$$

$$r_A = \frac{Cov(A_1, A_2)}{\sigma_A \sigma_A}$$

$$r_E = \frac{Cov(E_1, E_2)}{\sigma_{E_1}\sigma_{E_2}}$$





Genetic correlation – caused by - pleiotropic genes = genes with effect on both traits
- linkage – a gene that affects trait 1 is in LD with
a gene that affects trait 2

→ transient correlation – disappears with loss of LD

- quantifies the overall effect on both traits, across  $\underline{\mathsf{all}}$  loci

 $\rightarrow$  r<sub>A</sub> = 0 does <u>not</u> imply that there are <u>no</u> pleiotropic genes

**Environmental correlation** – caused by random environmental factors that affect both traits – measures the overall effect of <u>all</u> environmental factors

Some quantitative genetic math to show relationships among correlations:

$$Cov(P_1, P_2) = Cov(A_1 + E_1, A_2 + E_2) = Cov(A_1, A_2) + Cov(E_1, E_2)$$

$$\rightarrow r_P \sigma_{P_1} \sigma_{P_2} = r_A \sigma_{A_1} \sigma_{A_2} + r_E \sigma_{E_1} \sigma_{E_2}$$

$$ightharpoonup r_P \sigma_{P_1} \sigma_{P_2} = r_A h_1 \sigma_{P_1} h_2 \sigma_{P_2} + r_E e_1 \sigma_{P_1} e_2 \sigma_{P_2}$$
  $e^2 = 1 - h^2 = \text{prop.of phen.var. that is } \underline{\text{not}} \text{ add.genetic}$ 

$$\rightarrow r_P = r_A h_1 h_2 + r_F e_1 e_2$$

#### **GENETIC RELATIONSHIPS AND INBREEDING**

#### Are two alleles the same? Identity By State (IBS) versus Identity By Descent (IBD)

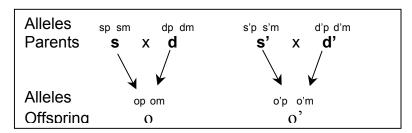
- IBS: if we can genotype individuals o and o' for this locus (QTL), then we can directly determine whether the alleles the two individuals carry are indeed the same

   if they are the same, this is referred to as the alleles being IBS.
- IBD: if we cannot genotype the locus (ie. the usual case), then we cannot determine IBS directly but, if o and o' have a common ancestor, than we can determine the probability that the two alleles are identical because they may have originated from a common ancestor

#### IBD probabilities from pedigree:

Prob(op is IBD to o'p) = P(op = o'p)

= probability that alleles op and o'p originated from the same allele of the common ancestor



Example IBD probabilities, coefficients of coancestry and additive and dominance coefficients

	IBD probabilities for pairs of alleles				Coancestry coefficient	Additive relationship coefficient	Dominance relationship coefficient
Individual o – o'	op-o'p	om-o'm	op-o'm	om-o'p	f <sub>oo</sub> ,	r <sub>oo</sub> ,	U <sub>oo</sub> ,
Sire(o) – Offspring(o')	1/2	0	0	1/2	1/4	1/2	0
Dam – Offspring	0	1/2	1/2	0	1/4	1/2	0
Paternal half-sibs	1/2	0	0	0	<sup>1</sup> / <sub>8</sub>	1/4	0
Full sibs	1/2	1/2	0	0	1/4	1/2	1/4
Identical twins	1	1	0	0	1/2	1	1

Coefficient of coancestry (also coeff. of kinship or consanguinity) between o and o' (See also Ch 5 p85)

=  $f_{oo'}$  = probability that an allele drawn at random from **o** is IBD to an allele drawn random from **o'** = average of the 4 possible IBD probabilities between alleles at o and o'

 $r_{\text{oo'}} = 2f_{\text{oo'}} = \text{coefficient of relationship} = \text{additive genetic relationship coefficient}$ 

**NOTE:**  $f_{oo}$  is also equal to the **coefficient of inbreeding** of a progeny produced by **o** and **o'** 

= probability that an individual's alleles are IBD

## From IBD probabilities to covariances of between relatives:

$$Cov(G_o,G_{o'}) = r_{oo'}V_A + u_{oo'}V_D$$

This equation applies to each locus that affects the trait but also to total genetic value; summing variances over loci ( $\mathbf{V}_{\mathbf{A}} = \Sigma \mathbf{V}_{\mathbf{A}i}$ ), this equation also applies to multiple loci

Thus, the genetic covariance (resemblance) between relatives is a function of their genetic relationships and genetic variance components.

- the additive genetic cov. between relatives = genetic relationship x add. genetic var.= roo'VA
  - relatives 'share' a portion  $\mathbf{r}_{oo}$ ' of their additive genetic variance because they share a portion  $\mathbf{r}_{oo}$ ' of their alleles
- the dominance genetic cov. betw. relatives = dom. relationship x dom. genetic var. =  $u_{oo'}V_D$

#### ALTERNATIVE DERIVATION OF ADDITIVE covariances based on quantitative genetics algebra

Model of phenotype: P = A + E E includes D, I, environment

Offspring phenotype:  $P_0 = A_0 + E_0 = \frac{1}{2}A_s + \frac{1}{2}A_d + RA_s + RA_d + E_0$ 



1/2 \* breeding value of parents

 $\triangleright$  Breeding value = 2\*E(P<sub>O</sub>-M) (by definition)

Includes some dominance and epistatic effects

RA<sub>s</sub>, RA<sub>d</sub> = random assortment / Mendelian sampling terms

- sampling of 1 of 2 parent alleles at each locus during meiosis
- by definition independent from other terms: Cov(A<sub>s</sub>,RA<sub>s</sub>) = 0

<u>Without</u> inbreeding:  $Var(RA_s) = \frac{1}{4}V_A$  Var(RA<sub>d</sub>) =  $\frac{1}{4}V_A$  (see derivation below)

With inbreeding:  $Var(RA_s) = \frac{1}{4}(1-F_s)V_A$   $Var(RA_d) = \frac{1}{4}(1-F_d)V_A$ 

Thus:  $Var(A_0) = Var(\frac{1}{2}A_s + \frac{1}{2}A_d + RA_s + RA_d) =$ 

=  $\frac{1}{4}V_A + \frac{1}{4}V_A + \frac{1}{4}V_A + \frac{1}{4}V_A = V_A$  (no inbreeding or selection)

#### Single locus derivation of Var(RA)

	Parent Geno- type	Fre- quen- cy	Genotypic value of parent [α=a+(q-p)d]		Offspring mean phenotype = ½*breeding	Trans- mitted allele	Fre- quen- cy	Offspring mean phenotype	Mendelian sampling term (RA)
-	A <sub>1</sub> A <sub>1</sub>	p <sup>2</sup>	а	$\frac{[\alpha-a+(q-p)a]}{2q(\alpha-qd)}$	value parent	A <sub>1</sub>	1	. ,	0
-		·	$d \qquad \qquad 2q(\alpha - qu)$ $d \qquad (q-p)\alpha + 2qd$		qα	A <sub>1</sub>	1/2	$egin{array}{c} oldsymbol{q}lpha \ oldsymbol{q}lpha \end{array}$	½α
А	A <sub>1</sub> A <sub>2</sub>	1A <sub>2</sub> 2pq		$\frac{1}{2}(q-p)\alpha$	A <sub>2</sub>	1/2	- <b>ρ</b> α	-½α	
	$A_2A_2$	$q^2$	-а	-2 $p(\alpha$ + $pd)$	<b>-ρ</b> α	$A_2$	1	<b>-ρ</b> α	0

E(RA<sub>s</sub>) = 
$$p^2(0) + 2pq\frac{1}{2}(\frac{1}{2}\alpha) + 2pq\frac{1}{2}(-\frac{1}{2}\alpha) + q^2(0) = 0$$

Without inbreeding: Var(RA<sub>s</sub>) =  $p^2(0)^2 + 2pq\frac{1}{2}(\frac{1}{2}\alpha)^2 + 2pq\frac{1}{2}(-\frac{1}{2}\alpha)^2 + q^2(0)^2$ =  $\frac{1}{2}pq\alpha^2 = \frac{1}{4}V_A$  ( $V_A = 2pq\alpha^2$ )

With inbreeding:  $F_s = Pr(two alleles in s <u>are</u> ibd) <math>\rightarrow RA_s = 0$ 

1- $F_s$  = Pr(two alleles in s not ibd)  $\rightarrow$  RA<sub>s</sub> = as in Table above

 $Var(RA_s) = F_s(0) + (1-F_s)\frac{1}{4}V_A = \frac{1}{4}(1-F_s)V_A$