

# Estimating social genetic effects

# Mixed model

- Assumed “true” model: 
$$P_i = A_{D,i} + E_{D,i} + \sum_{i \neq j} A_{S,j} + \sum_{i \neq j} E_{S,j}$$
- Mixed animal model: 
$$\mathbf{y} = \mathbf{X}\mathbf{b} + \mathbf{Z}_D\mathbf{a}_D + \mathbf{Z}_S\mathbf{a}_S + \mathbf{e}$$
  - $\mathbf{Z}_D\mathbf{a}_D$  = direct genetic effects of self
  - $\mathbf{Z}_S\mathbf{a}_S$  = social genetic effects of group members

## Example

- Mortality due to cannibalism in chickens
- 4 chickens per cage
- $\mathbf{Z}$ -matrices for two cages



$$\mathbf{Z}_D = \begin{bmatrix} 1 & 0 & 0 & 0 & 0 & 0 & 0 & 0 \\ 0 & 1 & 0 & 0 & 0 & 0 & 0 & 0 \\ 0 & 0 & 1 & 0 & 0 & 0 & 0 & 0 \\ 0 & 0 & 0 & 1 & 0 & 0 & 0 & 0 \\ 0 & 0 & 0 & 0 & 1 & 0 & 0 & 0 \\ 0 & 0 & 0 & 0 & 0 & 1 & 0 & 0 \\ 0 & 0 & 0 & 0 & 0 & 0 & 1 & 0 \\ 0 & 0 & 0 & 0 & 0 & 0 & 0 & 1 \end{bmatrix}$$

$$\mathbf{Z}_S = \begin{bmatrix} 0 & 1 & 1 & 1 & 0 & 0 & 0 & 0 \\ 1 & 0 & 1 & 1 & 0 & 0 & 0 & 0 \\ 1 & 1 & 0 & 1 & 0 & 0 & 0 & 0 \\ 1 & 1 & 1 & 0 & 0 & 0 & 0 & 0 \\ 0 & 0 & 0 & 0 & 0 & 1 & 1 & 1 \\ 0 & 0 & 0 & 0 & 1 & 0 & 1 & 1 \\ 0 & 0 & 0 & 0 & 1 & 1 & 0 & 1 \\ 0 & 0 & 0 & 0 & 1 & 1 & 1 & 0 \end{bmatrix}$$

# Mixed model

Example with 4 individuals in a group

$$\mathbf{y} = \mathbf{X}\mathbf{b} + \mathbf{Z}_D\mathbf{a}_D + \mathbf{Z}_S\mathbf{a}_S + \mathbf{e}$$

$$\begin{bmatrix} y_1 \\ y_2 \\ y_3 \\ y_4 \end{bmatrix} = \begin{bmatrix} 1 \\ 1 \\ 1 \\ 1 \end{bmatrix} [\mu] + \begin{bmatrix} 1 & 0 & 0 & 0 \\ 0 & 1 & 0 & 0 \\ 0 & 0 & 1 & 0 \\ 0 & 0 & 0 & 1 \end{bmatrix} \begin{bmatrix} A_{D,1} \\ A_{D,2} \\ A_{D,3} \\ A_{D,4} \end{bmatrix} + \begin{bmatrix} 0 & 1 & 1 & 1 \\ 1 & 0 & 1 & 1 \\ 1 & 1 & 0 & 1 \\ 1 & 1 & 1 & 0 \end{bmatrix} \begin{bmatrix} A_{S,1} \\ A_{S,2} \\ A_{S,3} \\ A_{S,4} \end{bmatrix} + \begin{bmatrix} e_1 \\ e_2 \\ e_3 \\ e_4 \end{bmatrix}$$

The residual summarizes both the direct and social non-genetic effects

$$\begin{bmatrix} e_1 \\ e_2 \\ e_3 \\ e_4 \end{bmatrix} = \begin{bmatrix} E_{D,1} + E_{S,2} + E_{S,3} + E_{S,4} \\ E_{D,2} + E_{S,1} + E_{S,3} + E_{S,4} \\ E_{D,3} + E_{S,1} + E_{S,2} + E_{S,4} \\ E_{D,4} + E_{S,1} + E_{S,2} + E_{S,3} \end{bmatrix}$$

# Mixed model: ASREML

- How to fit social effects into AsReml?
  - Use the “and()” statement in the model line
  - “and()” adds-up the Z-matrices

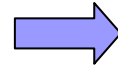
```
Data file groupsel.dat
self, m1, m2, m3, pheno
1 2 3 4 10
2 1 3 4 12
3 1 2 4 9
4 1 2 3 10
5 6 7 8 13
6 5 7 8 9
```

```
Pedigree file groupsel.ped
self sire dam
1 S1 D1
2 S2 D2
3 S3 D3
4 S4 D4
5 S5 D5
6 S6 D6
```

```
group selection
self !P
mate1 !P
mate2 !P
mate3 !P
ptype 1
groupsel.ped !MAKE !SKIP 2
groupsel.dat !maxiter=20 !SKIP 2
ptype ~ mu !r self mate1 and(mate2) and(mate3)
0 0 1
self 2
2 0 US !GP !+3
.1
0.01 .1
self 0 AINV
```

# Mixed model: residual variance structure

$$\begin{bmatrix} e_1 \\ e_2 \\ e_3 \\ e_4 \end{bmatrix} = \begin{bmatrix} E_{D,1} + E_{S,2} + E_{S,3} + E_{S,4} \\ E_{D,2} + E_{S,1} + E_{S,3} + E_{S,4} \\ E_{D,3} + E_{S,1} + E_{S,2} + E_{S,4} \\ E_{D,4} + E_{S,1} + E_{S,2} + E_{S,3} \end{bmatrix}$$



Can we simply fit a single residual?



What is the variance-covariance structure that emerges for the residual?

$$\begin{aligned} \text{Var}(e_i) &= \text{Var}(E_{D,i} + E_{S,j} + E_{S,k} + E_{S,l}) \\ &= \sigma_{E_D}^2 + 3\sigma_{E_S}^2 = \sigma_{E_D}^2 + (n-1)\sigma_{E_S}^2 \end{aligned}$$

$$\begin{aligned} \text{Cov}(e_i, e_j)_{\text{within\_grp}} &= \text{Cov}(E_{D,i} + E_{S,j} + E_{S,k} + E_{S,l}; E_{D,j} + E_{S,i} + E_{S,k} + E_{S,l}) \\ &= \text{Cov}(E_{D,i}, E_{S,i}) + \text{Cov}(E_{S,j}, E_{D,j}) + \text{Cov}(E_{S,k}, E_{S,k}) + \text{Cov}(E_{S,k}, E_{S,k}) \\ &= 2\sigma_{E_{DS}} + 2\sigma_{E_S}^2 = 2\sigma_{E_{DS}} + (n-2)\sigma_{E_S}^2 \end{aligned}$$

Two individuals have (n-2) group members in common → hence the (n-2)Var( $E_S$ )

Cov( $e_i, e_j$ ) = 0 between groups



# Mixed model: residual variance structure

- Within group, residuals are correlated
- There exist three biological VC
  - $\text{Var}(E_D)$
  - $\text{Cov}(E_D, E_S)$
  - $\text{Var}(E_S)$
- Statistically, we find only two VC
  - $\text{Var}(e)$
  - $\text{Cov}(e_i, e_j)_{\text{within\_grp}}$
- Hence, we cannot uniquely estimate all three biological VC
- $\text{Cov}(e_i, e_j)_{\text{within\_grp}} = 2\text{Cov}(E_D, E_S) + (n-2)\text{Var}(E_S)$ 
  - This can be either negative or positive
  - Probably positive in large groups
- Account for  $\text{Cov}(e_i, e_j)_{\text{within\_grp}} \rightarrow$  allow for correlated residuals

## ■ Residual variance structure

Two groups of 4 individuals each

$$\text{Var}(\mathbf{e}) = \mathbf{R}\sigma_e^2$$

with  $R_{ii} = 1$

$R_{ij} = \rho$  when  $i$  and  $j$  are group members

$R_{ij} = 0$  when  $i$  and  $j$  are in different groups

and

$$\sigma_e^2 = \sigma_{E_D}^2 + (n-1)\sigma_{E_S}^2$$

$$\rho = [2\sigma_{E_{DS}}^2 + (n-2)\sigma_{E_S}^2] / [\sigma_{E_D}^2 + (n-1)\sigma_{E_S}^2]$$

$$\mathbf{R} = \begin{bmatrix} 1 & \rho & \rho & \rho & 0 & 0 & 0 & 0 \\ \rho & 1 & \rho & \rho & 0 & 0 & 0 & 0 \\ \rho & \rho & 1 & \rho & 0 & 0 & 0 & 0 \\ \rho & \rho & \rho & 1 & 0 & 0 & 0 & 0 \\ 0 & 0 & 0 & 0 & 1 & \rho & \rho & \rho \\ 0 & 0 & 0 & 0 & \rho & 1 & \rho & \rho \\ 0 & 0 & 0 & 0 & \rho & \rho & 1 & \rho \\ 0 & 0 & 0 & 0 & \rho & \rho & \rho & 1 \end{bmatrix}$$

# Residual variance structure in ASREML

- Use the CORU statement in the R-structure definition
  - The starting value refers to  $\rho$

```
Data file groupsel.dat
group,nr,self,m1,m2,m3,phero
1 1 2 3 4 10
1 2 2 1 3 4 12
1 3 3 1 2 4 9
1 4 4 1 2 3 10
2 1 5 6 7 8 13
2 2 6 5 7 8 9
2 3 7 5 5 8 12
2 4 8 5 5 7 14
```

Include group in the data file,  
and a consecutive nr within  
the group

```
group selection
grp !I 2500
nr !I 4
self !P
mate1 !P
mate2 !P
mate3 !P
ptype 1
groupsel.ped !MAKE !SKIP 2
groupsel.dat !maxiter=20 !SKIP 2
ptype ~ mu !r self mate1 and(mate2) and(mate3)
1 2 1
2500 grp ID
4 nr CORU 0.1
self 2
2 0 US !GP !+3
.1
0.01 .1
self 0 AINV
```

Drawback: correlated residuals are computationally demanding and may converge slow



# Residual variance structure for large n

- $\rho = [2\text{Cov}(E_D, E_S) + (n-2)\text{Var}(E_S)] / \text{Var}(e)$ 
  - This is likely to be positive for large n
  - “Group members are similar” → you can fit a random group effect instead
  - This yields a simpler but equivalent model as long as  $\rho > 0$ .

$$\mathbf{y} = \mathbf{X}\mathbf{b} + \mathbf{Z}_D\mathbf{a}_D + \mathbf{Z}_S\mathbf{A}_S + \mathbf{Z}_g\mathbf{g} + \mathbf{e}^*$$

$$\begin{bmatrix} y_1 \\ y_2 \\ y_3 \\ y_4 \\ y_5 \\ y_6 \\ y_7 \\ y_8 \end{bmatrix} = \mathbf{X}\mathbf{b} + \mathbf{Z}_D\mathbf{a}_D + \mathbf{Z}_S\mathbf{A}_S + \begin{bmatrix} 1 & 0 \\ 1 & 0 \\ 1 & 0 \\ 1 & 0 \\ 0 & 1 \\ 0 & 1 \\ 0 & 1 \\ 0 & 1 \end{bmatrix} \begin{bmatrix} g_1 \\ g_2 \end{bmatrix} + \begin{bmatrix} e_1^* \\ e_2^* \\ e_3^* \\ e_4^* \\ e_5^* \\ e_6^* \\ e_7^* \\ e_8^* \end{bmatrix}$$

$$\text{Cov}(y_i, y_j | A) = 2\sigma_{E_{DS}} + (n-2)\sigma_{E_S}^2$$

$$\Rightarrow \sigma_g^2 = 2\sigma_{E_{DS}} + (n-2)\sigma_{E_S}^2$$

$$\sigma_g^2 + \sigma_{e^*}^2 = \sigma_e^2$$

$$\Rightarrow \sigma_{e^*}^2 = \sigma_e^2 - \sigma_g^2 = \sigma_{E_D}^2 - 2\sigma_{E_{DS}} + \sigma_{E_S}^2$$

Note: this redefines the residual and its variance → comparison of studies

$$\text{Var}(\mathbf{e}^*) = \mathbf{I}\sigma_{e^*}^2$$

Problem: This is valid only when  $\rho > 0$   
How do you know beforehand that  $\rho > 0$ ?



# Ignoring non-genetic social effects

What happens if you ignore  $E_S$ ?

- Simply fit  $\mathbf{y} = \mathbf{Xb} + \mathbf{Z}_D\mathbf{a}_D + \mathbf{Z}_S\mathbf{a}_S + \mathbf{e}$  with  $\text{Var}(\mathbf{e}) = \mathbf{I}$   $\text{Var}(\mathbf{e})$
- This assumes that  $\rho = 2\text{Cov}(E_D, E_S) + (n-2)\text{Var}(E_S) = 0$
- Either: social effects are assumed fully heritable,  $\text{Var}(E_S) = 0$
- or  $(n-2)\text{Var}(E_S) = -2\text{Cov}(E_D, E_S)$
- These are very strong a priori assumptions
- This is not an issue of statistical significance or not, always allow for  $E_S$

## Consequences of ignoring $E_S$

- $\text{Var}(E_S)$  ends up in  $\text{Var}(A_S) \rightarrow$ 
  - Severe overestimation of (social) genetic variance
- Bijma et al., 2007b
  - Estimated  $\rho = 0.09$  ( $P < 0.001$ )
  - Using  $\text{Var}(\mathbf{e}) = \mathbf{I}$   $\text{Var}(\mathbf{e}) \rightarrow \text{Var}(\text{TBV})$  overestimated by a factor of 2.6!

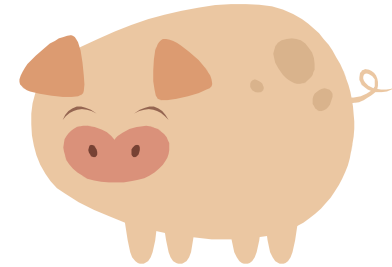


# Message

Estimated genetic parameters for social effects are extremely sensitive to what other components you fit in the model

**Model selection is a key issue**

# Ignoring non-genetic social effects



- Ignoring social effects may bias estimation of *classical* heritability when group members are related
- Feed intake pigs,  $n = 8$ , Bergsma et al 2008
- Average relatedness within group,  $r = 0.18$
- Classical model  $y = Xb + Za + e$ 
  - Estimated  $h^2 = 0.41$
- Accounting for group effect  $y = Xb + Za + Z_g g + e$ 
  - Estimated  $h^2 = 0.18$
  - Physically pens were identical
- Due to social effects and large  $n$ , group members are similar ( $\rho > 0$ )
  - Similar group members  $\rightarrow$  similar relatives  $\rightarrow h^2 \uparrow$



# Statistical models for socially affected traits

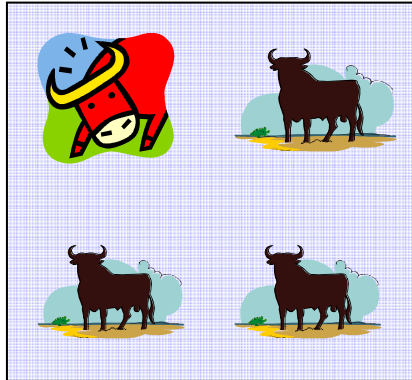
- Which fixed and random effect to include?
  - The social effect is “phenotypic”
    - It may have fixed, random, and genetic components
  - Not all biological components may be estimable (e.g.  $\text{Var}(E_S)$ ,  $\text{Cov}(E_D, E_S)$  and  $\text{Var}(E_D)$ )
    - Derive the resulting variance structure within and between groups
  - Statistical significance of correction factors is not the primary issue
    - We know that  $h^2 \neq 100\%$ , account for E, also when  $p > 0.1$



# Statistical models for socially affected traits

- The social effect is “phenotypic”
  - Include fixed effects for the group member
    - $Y = \{X_D b_D + Z_D a_D + e_D\} + \{X_S b_S + Z_S a_S + e_S\}$
    - Sex, age or breed of the group member
      - These are usually easy to estimate
  - Include random effects for the group member
    - $Y = \{X_D b_D + Z_D a_D + .. + e_D\} + \{X_S b_S + Z_S a_S + .. + e_S\}$
    - Litter of the group member (non-genetic maternal social effect)
    - Permanent effects (repeated records)
    - Mother of the group member (genetic maternal social effect)
      - These are not always easy to estimate
      - Test sensitivity of your social VC for other model components
  - Derive the theoretically expected (residual) variance structure
    - And allow for it in your statistical model

# Example: mixed breeds in beef cattle



- Allow for a social fixed effect of the breed of the group members
  - Angus or Hereford (A,H)

$$\mathbf{y} = \mathbf{X}_D \mathbf{b}_D + \mathbf{X}_S \mathbf{b}_S + \mathbf{Z}_D \mathbf{a}_D + \mathbf{Z}_S \mathbf{a}_S + \mathbf{e}$$

$$\begin{bmatrix} y_1 \\ y_2 \\ y_3 \\ y_4 \end{bmatrix} = \begin{bmatrix} 1 & 0 \\ 0 & 1 \\ 0 & 1 \\ 0 & 1 \end{bmatrix} \begin{bmatrix} H_D \\ A_D \end{bmatrix} + \begin{bmatrix} 0 & 3 \\ 1 & 2 \\ 1 & 2 \\ 1 & 2 \end{bmatrix} \begin{bmatrix} H_S \\ A_S \end{bmatrix} + \mathbf{Z}_D \mathbf{a}_D + \mathbf{Z}_S \mathbf{a}_S + \mathbf{e}$$

Breed of self

Breed of group members

Use the “and()” statement in Asreml



# Estimability of social genetic VC

- Little research has been done
- So far
  - Relatedness within and between groups is critical
  - FS-groups is impossible, irrespective of pedigree
    - You cannot distinguish direct from social effects
  - HS-groups without FS in the data is also impossible
  - Be carefull with structuring families across groups (e.g. Wolf PNAS paper → see Bijma et al., 2007b)
  - Random groups (with respect to relatedness) works well
    - But is probably not optimal
  - Combining two families per group is an option
  - This may be useful when tagging is difficult (marine species?)
  - You cannot fit a **fixed** group effect
    - Problematic???
    - Avoid confounding of physically good pens with certain families
- Data requirements
  - ~4 times more than for direct effect only (random groups)
  - More if groups are larger



# Dealing with variable group size

## ■ Case 1

- Underlying parameters do not depend on group size (no “true” G x group-size interaction)
- Genetic VC are constant [in particular:  $V(A_S) \neq f(n)$ ]
- Issue is impact of  $n$  on non-genetic section of model

### Correlated residual model

$$\mathbf{y} = \mathbf{X}\mathbf{b} + \mathbf{Z}_D\mathbf{a}_D + \mathbf{Z}_S\mathbf{a}_S + \mathbf{e}$$

$$\text{Var}(\mathbf{e}) = \mathbf{R}\sigma_e^2, R_{ii} = 1, R_{ij, \text{within}} = \rho$$

$$\rho = [2\sigma_{E_{DS}}^2 + (n-2)\sigma_{E_S}^2] / \sigma_e^2$$

$$\sigma_e^2 = \sigma_{E_D}^2 + (n-1)\sigma_{E_S}^2$$

$\rho$  and  $\text{Var}(\mathbf{e})$  depend on group size



Fit heterogeneous  $\text{Var}(\mathbf{e})$  and  $\rho$

Should be possible in AsReml

# Dealing with variable group size

## Random group effect model

$$\mathbf{y} = \mathbf{X}\mathbf{b} + \mathbf{Z}_D\mathbf{a}_D + \mathbf{Z}_S\mathbf{a}_S + \mathbf{Z}_g\mathbf{g} + \mathbf{e}^*$$

$$\text{Var}(\mathbf{e}^*) = \mathbf{I}\sigma_{e^*}^2, \sigma_{e^*}^2 = \sigma_{E_D}^2 - 2\sigma_{E_{DS}} + \sigma_{E_S}^2$$

$$\text{Var}(\mathbf{g}) = [2\sigma_{E_{DS}} + (n-2)\sigma_{E_S}^2]$$

Var(e) does not depend on group size

but Var(g) does



Fit heterogeneous Var(g)

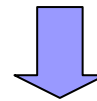
Consequences of ignoring heterogeneity of variance have not been investigated

The genetic term is “automatically” heterogeneous because the number of group members in  $\mathbf{Z}_S$  varies → If non-genetic heterogeneity of variance exists, you may expect it to end up in (and thus inflate) the genetic terms

# Dealing with variable group size

## ■ Case 2

- Variance of social effects depend on group size,  
 $\text{Var}(A_S) = f(n)$



Consider:  $\sigma_{TBV}^2 = \sigma_{A_D}^2 + 2(n-1)\sigma_{A_{DS}} + (n-1)^2 \sigma_{A_S}^2$

- In large groups, heritable variance is very large → this may not make sense
- The social effects per individual must become smaller →  $\text{Var}(A_S)$  must go down with  $n$
- This is not true GxE-interaction, just scaling or “dilution”
- i.e.  $\text{Corr}(A_{S,i,n=4}, A_{S,i,n=5}) = 1$ , but  $\text{Var}(A_{S,i,n=4}) > \text{Var}(A_{S,i,n=5})$
- We found such results for growth in pigs



# Dealing with variable group size

- Accounting for decreasing  $\text{Var}(A_S)$  with  $n$ 
  - Diluting social effects depending on  $n - 1$

$$y_i = \text{fixed} + A_{D,i} + c_{n-1} \sum_{n-1} A_{S,j} + e_i$$

- $c$  is an unknown constant depending on  $(n - 1)$  (Arango et al., 2005, JAS)
- $c = 1 \rightarrow \text{Var}(A_S)$  is independent of  $n \rightarrow \text{Var}(\text{TBV})$  increases with  $n$
- $c = 1/(n - 1) \rightarrow$  the sum of social effects is constant
  - $\rightarrow \text{Var}(\text{TBV})$  is independent of  $n$
- find best  $c$  iteratively using AsReml
  - E.g.  $c = (n-1)^x$ , vary  $x$  from 0 to  $-1$
  - This is like random regression, slope = age \*  $b_{\text{slope}}$ 
    - but “age” is unknown  $\rightarrow$  iterate to ReML value



# An alternative model useful for BVE (Abe Huisman)

- Interesting when:
  - If your BVE-software does not allow for social effects
  - Your data file becomes too large when you add all group members
- Idea
  - Direct effects are expressed in self
  - Social effects are expressed in group members
  - → use conventional bivariate analysis with two traits:
    1. Own performance
    2. Mean performance of group members
  - This fits in ordinary BVE software
- Issues
  - This model does not properly fit when group members are related
  - → not at all robust for VCE, seems less important for BVE
  - This has not been extensively tested!



# Summary on VCE

- Social variance components can be estimated
  
- They are sensitive to BIAS
  - Confounding with non-genetic social effects
  - Fit fixed effects also for social component
  - Sensitivity analysis is important
  - Think of the biological interpretation of your model
  
- More research is needed on
  - Optimum designs for analysis (relatedness)
  - Varying group size
  - .....